

PCA & SVD

Note Title

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Recall the **centered data matrix**
 $\tilde{X} := [\tilde{x}_1 \cdots \tilde{x}_n] \in \mathbb{R}^{d \times n}$

$$\tilde{x}_j := x_j - \bar{x}, \quad \bar{x} := \frac{1}{n} \sum_{i=1}^n x_i,$$

and the **sample covariance matrix**

$$S := \frac{1}{n} \tilde{X} \tilde{X}^T$$

Then, **PCA** is nothing but the **eigendecomposition of S**

$$S = \Phi \Lambda \Phi^T, \quad \Lambda = \text{diag}(\lambda_1, \dots, \lambda_d)$$

$$\lambda_1 \geq \lambda_2 \geq \dots \geq \lambda_d \geq 0.$$

$\Phi := [\phi_1 \cdots \phi_d] \in \mathbb{R}^{d \times d}$ is an ortho. matrix, and $\{\phi_1, \dots, \phi_d\}$ form an ONB of $\mathbb{R}^d$.

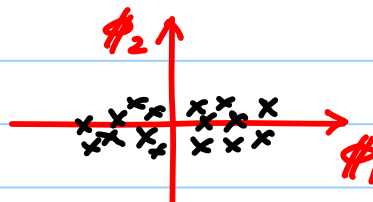
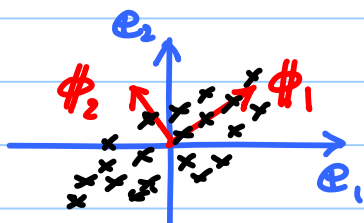
$\phi_j^T \tilde{X}$ is said to be the **j th**

principal components of $\tilde{X}$.

These are nothing but the expansion coefficients of $\tilde{X}$ w.r.t. the ONB vector ϕ_j .

If $\tilde{X}$ forms a "cigar" shape,

then $\phi_j^T \tilde{X}$ are the coordinate values of $\tilde{X}$ under the rotated axes



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- Hence viewing the given dataset under the principal axes $\Phi_1, \Phi_2, \dots$, provides us better interpretations of the data than viewing them under the original axes $\mathcal{E}_1, \mathcal{E}_2, \dots$.
- PCA is also often used as a tool to do **dimension reduction** and **feature extraction** by keeping only **top k** PCA coordinates where $k \ll d$, i.e.,

$$\Phi_k := [\Phi_1 \dots \Phi_k] \in \mathbb{R}^{d \times k}$$

$$\mathbb{R}^d \ni \tilde{\mathbf{x}}_j \mapsto \underbrace{\Phi_k^T \tilde{\mathbf{x}}_j}_{\text{top } k \text{ PCA coordinates or top } k \text{ Principal components of } \tilde{\mathbf{x}}_j} \in \mathbb{R}^k$$

top k PCA coordinates
or top k Principal
components of $\tilde{\mathbf{x}}_j$.

Note that using these **top k** principal components, we can approximate the original data $\mathbf{x}_j$ by

$$\mathbf{x}_j \approx \bar{\mathbf{x}} + \Phi_k \Phi_k^T \tilde{\mathbf{x}}_j$$

Of course the approximation gets better and better as k increases.

In fact, if $k = d$, then $\mathbf{x}_j$ is recovered exactly (within machine ϵ).

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Now we'll face the problem when we compute the eigendecomposition of $S = \Phi \Lambda \Phi^T$:

- (1) If d is large, we cannot compute this eigendecomposition because we cannot hold $\Phi \in \mathbb{R}^{d \times d}$ in computer memory, and its computational cost is $O(d^3)$, i.e., too expensive to compute.
- (2) Fortunately, we often do not need all d eigenvectors, most likely, only first k eigenvectors $k \ll d$.
- (3) Moreover if $d > n$, then $\text{rank}(S) = n - 1$ if $\mathbf{x}_j$'s are linearly indep. So, after the first $n - 1$ eigenvectors are useless!

Why? $S = \frac{1}{n} \tilde{X} \tilde{X}^T = \frac{1}{n} \{ \underbrace{\tilde{\mathbf{x}}_1 \tilde{\mathbf{x}}_1^T}_{\text{rank 1}} + \dots + \underbrace{\tilde{\mathbf{x}}_n \tilde{\mathbf{x}}_n^T}_{\text{rank 1}} \}$

So looks like $\text{rank}(S) = n$.

But since $\tilde{\mathbf{x}}_1 + \dots + \tilde{\mathbf{x}}_n = \mathbf{0}$ because the mean $\bar{\mathbf{x}}$ is subtracted from each data vector $\mathbf{x}_j$ (i.e., $\tilde{\mathbf{x}}_j = \mathbf{x}_j - \bar{\mathbf{x}}$)
Hence, S loses 1 rank.

So, $\text{rank}(S) = n - 1$.

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Now, let's consider the **reduced** SVD of $\tilde{X}$:

$$\tilde{X} = \hat{U} \hat{\Sigma} V^T$$

$$\begin{array}{c} d \geq n \\ \boxed{\tilde{X}} = \boxed{\hat{U}} \boxed{\hat{\Sigma}} \boxed{V^T} \end{array} \quad \begin{array}{c} d < n \\ \boxed{\tilde{X}} = \boxed{U} \boxed{\hat{\Sigma}} \boxed{\hat{V}^T} \end{array}$$

Just consider the "neo-classical" setting, i.e., $d \geq n$ (e.g., the face image database)

Then consider the sample covariance matrix S using the above SVD:

$$\begin{aligned} S &= \frac{1}{n} \tilde{X} \tilde{X}^T = \frac{1}{n} \hat{U} \hat{\Sigma} \underbrace{V^T V}_{=I} \hat{\Sigma}^T \hat{U}^T \\ &= \frac{1}{n} \hat{U} \hat{\Sigma} \hat{\Sigma}^T \hat{U}^T = \frac{1}{n} \hat{U} \hat{\Sigma}^2 \hat{U}^T \end{aligned}$$

Now $\hat{\Sigma} = \text{diag}(\sigma_1, \dots, \sigma_{n-1}, \underline{\underline{0}})$
if $x_1, \dots, x_n$ are linearly indep.

$$\text{So, } \hat{\Sigma}^2 = \text{diag}(\sigma_1^2, \dots, \sigma_{n-1}^2, 0).$$

Finally, S can be written as

$$S = \underbrace{\hat{U}}_{\substack{\uparrow \\ \text{columns are orthonormal}}} \underbrace{\left(\frac{1}{n} \hat{\Sigma}^2 \right)}_{= \text{diag}(\sigma_1^2/n, \dots, \sigma_{n-1}^2/n, 0)} \hat{U}^T$$

Comparing this with the eigendecomposition

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$S = \Phi \Lambda \Phi^T$, we can conclude that

$$\begin{cases} \Phi[:, 1:n] = \hat{U} \\ \Lambda[1:n, 1:n] = \frac{1}{n} \hat{\Sigma}^2 = \text{diag}(\sigma_1^2/n, \dots, \sigma_{n-1}^2/n, 0) \end{cases}$$

In fact, only the $1:n-1$ portion is useful since $\sigma_n = 0$.

Hence, we should **use the reduced SVD of $\tilde{X}$ (not S) for computing PCA!!**
Do not use the eigendecomposition of S unless d is small.

Note: $\tilde{X} V = \hat{U} \hat{\Sigma} V^T V = \hat{U} \hat{\Sigma}$
 $= [\tilde{X} v_1, \dots, \tilde{X} v_n] = [\sigma_1 u_1, \dots, \sigma_{n-1} u_{n-1}, 0]$
 So, $u_j = \frac{1}{\sigma_j} \tilde{X} v_j$, $j = 1, \dots, n-1$.

In other words, **each principal axis u_j is just a linear combination of the (centered) input vectors $\tilde{X}_1, \dots, \tilde{X}_n$!**

Now let's do Julia experiments using the face image database consisting of 143 faces each of which has $128 \times 128 = 16384$ pixels, i.e., $d = 16384$, $n = 143$.