Conditional Distributions, Log-linear Models, and Disclosure Limitation Methods*

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An Example of Bounds for Table Entries

[These tables are taken from Dobra and Fienberg (2001). We include some additional tables based on a 10% random sample of the data.]

			B	no		yes	
F	E	D	C / A	no	yes	no	yes
neg	< 3	< 140	no	44	40	112	67
			yes	129	145	12	23
		≥ 140	no	35	12	80	33
			yes	109	67	7	9
	≥ 3	< 140	no	23	32	70	66
			yes	50	80	7	13
		≥ 140	no	24	25	73	57
			yes	51	63	7	16
pos	< 3	< 140	no	5	7	21	9
			yes	9	17	1	4
		≥ 140	no	4	3	11	8
			yes	14	17	5	2
	≥ 3	< 140	no	7	3	14	14
			yes	9	16	2	3
		≥ 140	no	4	0	13	11
			yes	5	14	4	4

Table 1: Prognostic factors in coronary heart disease. Source: Edwards and Havranek (1985).

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F	E	D	C	B / A	no / no	no / yes	yes / no	yes / yes
neg	< 3	< 140	no		[0,88]	[0,62]	[0,224]	[0,117]
			yes		[0,261]	[0,246]	[0,25]	[0,38]
		≥ 140	no		[0,88]	[0,62]	[0,224]	[0,117]
			yes		[0,261]	[0,151]	[0,25]	[0,38]
	≥ 3	< 140	no		[0,58]	[0,60]	[0,170]	[0,148]
			yes		[0,115]	[0,173]	[0,20]	[0,36]
		≥ 140	no		[0,58]	[0,60]	[0,170]	[0,148]
			yes		[0,115]	[0,173]	[0,20]	[0,36]
pos	< 3	< 140	no		[0,88]	[0,62]	[0,126]	[0,117]
			yes		[0,134]	[0,134]	[0,25]	[0,38]
		≥ 140	no		[0,88]	[0,62]	[0,126]	[0,117]
			yes		[0,134]	[0,134]	[0,25]	[0,38]
	≥ 3	< 140	no		[0,58]	[0,60]	[0,126]	[0,126]
			yes		[0,115]	[0,134]	[0,20]	[0,36]
		≥ 140	no		[0,58]	[0,60]	[0,126]	[0,126]
			yes		[0,115]	[0,134]	[0,20]	[0,36]

Table 2: Bounds for cell counts in the coronary heart disease table given margins corresponding to [BF][ADE][ABCE].

F	E	D	C	B / A	no / no	no / yes	yes / no	yes / yes
neg	< 3	< 140	no		7	6	14	4
			yes		6	19	2	3
		≥ 140	no		6	3	5	1
			yes		8	6	1	1
	≥ 3	< 140	no		0	6	8	15
			yes		4	9	0	1
		≥ 140	no		2	2	6	4
			yes		3	3	0	3
pos	< 3	< 140	no		0	1	2	0
			yes		1	1	0	0
		≥ 140	no		0	0	1	0
			yes		1	1	0	0
	≥ 3	< 140	no		2	0	2	2
			yes		1	0	2	1
		≥ 140	no		0	0	3	2
			yes		1	2	0	0

Table 3: 10% sample selected from the population with coronary heart disease.

F	E	D	C	B	no		yes	
				A	no	yes	no	yes
neg	< 3	< 140	no		[0,13]	[0,10]	[0,22]	[0,5]
			yes		[0,16]	[4,27]	[0,3]	[0,4]
		≥ 140	no		[0,13]	[0,10]	[0,22]	[0,5]
			yes		[0,16]	[0,12]	[0,3]	[0,4]
	≥ 3	< 140	no		[0,4]	[0,8]	[0,19]	[0,23]
			yes		[0,9]	[0,14]	[0,2]	[0,5]
		≥ 140	no		[0,4]	[0,8]	[0,15]	[0,16]
			yes		[0,9]	[0,14]	[0,2]	[0,5]
pos	< 3	< 140	no		[0,11]	[0,10]	[0,15]	[0,5]
			yes		[0,11]	[0,11]	[0,3]	[0,4]
		≥ 140	no		[0,11]	[0,10]	[0,15]	[0,5]
			yes		[0,11]	[0,11]	[0,3]	[0,4]
	≥ 3	< 140	no		[0,4]	[0,8]	[0,15]	[0,15]
			yes		[0,9]	[0,11]	[0,2]	[0,5]
		≥ 140	no		[0,4]	[0,8]	[0,15]	[0,15]
			yes		[0,9]	[0,11]	[0,2]	[0,5]

Table 4: Bounds for cell counts in the 10% sample table given margins corresponding to [BF][ADE][ABCE].

The Diaconis-Sturmfels Algorithm (1998)

[This material is extracted from Fienberg, Makov, Meyer, and Steele (2001).]

Let $\mathbf{n}$ is the observed table, μ is the table of expected values under the model, $\mathbf{c}$ is the constraint vector representing the conditioning involving marginal totals, and $S(\mathbf{c})$ is the set of all nonnegative tables satisfying the marginal constraints. Let $\{f_1, f_2, \ldots, f_L\}$ be a generating set for the tables in $S(\mathbf{c})$.

> *Lemma*: Let σ be a positive function on $S(\mathbf{c})$. Generate a Markov chain on $S(\mathbf{c})$ by choosing I uniformly in $\{1, 2, \ldots, L\}$ and $\epsilon = \pm 1$ with probability 1/2 independently of I. If the chain is currently at $\mathbf{m}$ it moves to $\mathbf{m}' = \mathbf{m} + \epsilon f_I$ (provided that $\mathbf{m}' \in S(\mathbf{c})$ with probability $\min(1, \sigma(\mathbf{m}')/\sigma(\mathbf{m}))$. In all other cases the chain stays at $\mathbf{m}$. This is a connected, reversible Markov chain on $S(\mathbf{c})$ with a stationary distribution proportional to $\sigma(\mathbf{m})$.

By decoupling the "positive" and "negative" versions of the move to f_i for $i = 1, 2, \ldots, L$, Diaconis and Sturmfels get transition probabilities that can be calculated for any model, even for nondecomposable loglinear models, as long as the margins we condition on are those that correspond to the minimal sufficient statistics. The argument is as follows.

From Haberman (1974), we know that the underlying hypergeometric distribution for the exact distribution of the table under a loglinear model given a set of marginal constraints is

$$\sigma(\mathbf{n}) = \frac{(\prod_{i \in I} \frac{1}{n(i)!}) exp[\mathbf{n}, \mu]}{\sum_{\mathbf{m} \in S(\mathbf{c})} (\prod_{i \in I} \frac{1}{m(i)!}) exp[\mathbf{m}, \mu]} \tag{1}$$

where $\mathbf{n}$ is the observed table, μ is the table of expected values under the model, $\mathbf{c}$ is the constraint vector representing the conditioning involving marginal totals, and $S(\mathbf{c})$ is the set of all nonnegative tables satisfying the marginal constraints. When we condition on the margins that correspond to the minimal sufficient statistics under the model, the probabilities in equation (1) simplify because all of the exponential components are the same, yielding:

$$\sigma(\mathbf{n}) = \frac{\prod_{i \in I} \frac{1}{n(i)!}}{\sum_{\mathbf{m} \in S(\mathbf{c})} (\prod_{i \in I} \frac{1}{m(i)!})}. \tag{2}$$

The denominator in equation (2) is the same for each table with the specified margins and so the ratio of two such probabilities is only a function of the corresponding numerators.

Fienberg, Makov and Steele (1998), and Fienberg, Makov, Meyer and Steele (2001) analyze the following example in Table 5, drawn from data from the 1990 U.S. decennial census.

Gender = Male

	Income Level			
Race	$\leq$ \$10,000	> \$10000 and $\leq$ \$25000	> \$25000	Total
White	96	72	161	329
Black	10	7	6	23
Chinese	1	1	2	4
Total	107	80	169	356

Gender = Female

	Income Level			
Race	$\leq$ \$10,000	> \$10000 and $\leq$ \$25000	> \$25000	Total
White	186	127	51	364
Black	11	7	3	21
Chinese	0	1	0	1
Total	197	135	54	386

Table 5: Three-way cross-classification of Gender, Race, and Income for a selected U.S. census tract. (*Source*: 1990 Census Public Use Microdata Files)

In Table 6 we present the maximum likelihood estimates for the expected counts corresponding to the entries in Table 5 under the no 2nd-order interaction model with multinomial sampling We computed these in S-plus. The likelihood ratio chi-squared value for the fit of this model was 2.89 on 4 d.f. This is indicative of a moderately good model fit, although it is actually somewhat difficult to assess the fit given the sparseness of the row in the first layer which has a total count of 1 in it.

Gender = Male

Race	Income Level $\leq \$10,000$	$> \$10000$ and $\leq \$25000$	$> \$25000$	Total
White	97.09	72.15	159.76	329
Black	9.21	6.41	7.38	23
Chinese	0.70	1.44	1.86	4
Total	107	80	169	356

Gender = Female

Race	Income Level $\leq \$10,000$	$> \$10000$ and $\leq \$25000$	$> \$25000$	Total
White	184.91	126.85	52.24	364
Black	11.79	7.58	1.62	21
Chinese	0.30	0.56	0.14	1
Total	197	135	54	386

Table 6: Maximum likelihood estimates for data in Table 5 under the no 2nd-order interaction model.

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