# Doob decomposition and martingales with bounded increments

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#### Theorem

Assume that  $X_n$ ,  $n \ge 0$  is a submartingale w.r.t.  $\mathcal{F}_n$ . Then  $X_n$  can be written in a unique way (up to a.s. equality) as  $X_n = M_n + A_n$ , where

- M<sub>n</sub> is a martingale; and
- ②  $A_n$  is increasing, predictable, and  $A_0 = 0$ .

Note that  $A_{\infty} = \lim_{n \to \infty} A_n$  exists (but it may be  $\infty$ ).

#### Proof.

Assuming the decomposition:

$$E[X_n \mid \mathcal{F}_{n-1}] = E[M_n \mid \mathcal{F}_{n-1}] + E[A_n \mid \mathcal{F}_{n-1}]$$

$$= M_{n-1} + A_n$$

$$= X_{n-1} + A_n - A_{n-1}$$
and so
$$A_n = \sum_{k=1}^n (E[X_k \mid \mathcal{F}_{k-1}] - X_{k-1}),$$

and so

and we have proved uniqueness. We need to check that so defined  $A_n$  are increasing, predictable, with  $A_0 = 0$  (all clear), and that  $X_n - A_n$  is a martingale:

$$E[X_n - A_n \mid \mathcal{F}_{n-1}] = E[X_n \mid \mathcal{F}_{n-1}] - A_n = X_{n-1} - A_{n-1}.$$

Now let  $X_n$  be a martingale with  $X_0 = 0$ ,  $EX_n^2 < \infty$  for all n. Then  $X_n^2$  is a submartingale, so by Doob,  $X_n^2 = M_n + A_n$ , where  $M_n$  is a martingale and

$$\langle X \rangle_n := A_n = \sum_{k=1}^n \left( E[X_k^2 \mid \mathcal{F}_{k-1}] - X_{k-1}^2 \right)$$

is called the *bracket*, or *quadratic variation*, process for  $X_n$ .

It is used to measure fluctuations in  $X_n$  so it has a similar role to variance of a r.v.

**Example**. Let  $\xi_i$  be independent with  $E\xi_i = 0$  and  $E\xi_i^2 < \infty$ ,  $S_n = \xi_1 + \cdots + \xi_n$ . Then

$$E[S_n^2 \mid \mathcal{F}_{n-1}] = E[S_{n-1}^2 + 2S_{n-1}\xi_n + \xi_n^2 \mid \mathcal{F}_{n-1}] = S_{n-1}^2 + E\xi_n^2,$$

so  $\langle S \rangle_n = \sum_{k=1}^n E \xi_k^2$  is deterministic in this case.

If N is a stopping time, then

$$X_{n\wedge N}^2 = M_{n\wedge N} + A_{n\wedge N}$$

and  $M_{n \wedge N}$  is a martingale, while  $A_{n \wedge N}$  is predictable, as

$$A_{n \wedge N} = \sum_{k=1}^{n-1} A_k 1_{\{N=k\}} + A_n 1_{\{N \geq n\}} = \sum_{k=1}^{n-1} A_k 1_{\{N=k\}} + A_n 1_{\{N \leq n-1\}^c}$$

It follows that the bracket of  $X_{n \wedge N}$  is  $\langle X \rangle_{n \wedge N}$ .

Assume until further notice that  $X_n$  is a martingale with  $X_0 = 0$ ,  $EX_n^2 < \infty$  for all n, and  $A_n = \langle X \rangle_n$ .

### **Proposition**

$$E[\sup_{k\geq 1} X_k^2] \leq 4EA_{\infty}.$$

#### Proof.

We first use the  $L^2$ -maximum inequality

$$E[\sup_{1 \le k \le n} X_k^2] \le 4EX_n^2 = 4EA_n \le 4EA_\infty$$

and then MCT.



### **Proposition**

On  $\{A_{\infty} < \infty\}$ ,  $\lim X_n$  exists and is finite a.s.

#### Proof.

Fix a>0 and let  $N=\inf\{n:A_{n+1}>a\}$ , which is a stopping time (by predictability of  $A_n$ ). Then  $X_{n\wedge N}^2=M_{n\wedge N}+A_{n\wedge N}$ , and so by the previous proposition,

$$E[\sup_{n} X_{n \wedge N}^{2}] \leq 4EA_{N} \leq 4a.$$

Therefore,  $\lim_{n\to\infty}X_{n\wedge N}$  exists a.s. and in  $L^2$ , and then  $\lim_{n\to\infty}X_n$  exists a.s. on  $\{N=\infty\}=\{A_\infty\leq a\}$ , and finally  $\lim_{n\to\infty}X_n$  exists a.s. on  $\{A_\infty<\infty\}=\cup_{a=1}^\infty\{A_\infty\leq a\}$ .

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#### Example. Let

$$\xi_n = \begin{cases} \pm 2^n & \text{w.p. } 1/2^{n+1} \\ 0 & \text{w.p. } 1 - 1/2^n \end{cases}$$

and assume these are independent and  $S_n = \xi_1 + \ldots + \xi_n$ . Then

$$\langle S \rangle_n = \sum_{k=1}^n 2^{2k} 2^{-k} = \sum_{k=1}^n 2^k \to \infty,$$

but  $S_n$  converges a.s. as  $P(\xi_n \neq 0 \text{ i.o.}) = 0$ .

So it is not in general true that  $X_n$  does not converge on  $\{A_{\infty} = \infty\}$ .

#### Theorem

Assume that  $X_n$  is a martingale and that there exists a (deterministic) constant M such that  $|X_{n+1}-X_n|\leq M$  for all  $n\geq 0$ . Let

$$C = \{ \lim_{n} X_n \text{ exists and is finite} \}$$

$$D = \{ \lim_{n} \sup_{n} X_n = \infty, \lim_{n} \inf_{n} X_n = -\infty \}$$

#### Then:

- (a)  $C = \{A_{\infty} < \infty\}$  a.s.; and
- (b)  $P(C \cup D) = 1$ .

#### Proof.

Assume WLOG that  $X_0 = 0$ .

To prove (b), define, for K>0, the stopping time  $N=N_K=\inf\{n:X_n\geq K\}$ . Then  $X_{n\wedge N}$  is a martingale and, by bounded increments,  $X_{n\wedge N}\leq K+M$ . Thus  $\lim X_{n\wedge N}$  exists a.s. and is finite. So,  $\lim X_n$  exists and is finite a.s. on

$$\cup_{K=1}^{\infty}\{\textit{N}_{K}=\infty\}=\{\sup\textit{X}_{\textit{n}}<\infty\}=\{\limsup\textit{X}_{\textit{n}}<\infty\}.$$

By symmetry,  $\lim X_n$  also exists and is finite a.s. on  $\{\lim \inf X_n > -\infty\}$ .

#### Proof, continued.

To prove (a), it is now equivalent to show (by (b) and the previous proposition) that

$$P(A_{\infty}=\infty,\sup |X_n|<\infty)=0.$$

Let now  $N = N_K = \inf\{n : |X_n| \ge K\}$ . We know that

 $E[X_{n\wedge N}^2 - A_{n\wedge N}] = 0$  and  $|X_{n\wedge N}| \leq K + M$ .

So,  $EA_{n\wedge N} \leq (K+M)^2$ .

By MCT,  $EA_N \leq (K+M)^2$ .

So on  $\{N = \infty\}$ ,  $A_{\infty} < \infty$  a.s., that is,

 $P(N_K = \infty, A_\infty = \infty) = 0$  and then

$$P(\cup_{K=1}^{\infty}\{N_K=\infty\},A_{\infty}=\infty)=0,$$

but 
$$\bigcup_{K=1}^{\infty} \{N_K = \infty\} = \{\sup |X_n| < \infty\}.$$



### Proposition (Borel-Cantelli part 3)

Let  $\mathcal{F}_n$  be a filtration and  $A_n \in \mathcal{F}_n$ ,  $n \ge 1$ . Then, a.s.,

$$\{A_n \text{ i.o.}\} = \left\{ \sum_{n=1}^{\infty} E[1_{A_n} \mid \mathcal{F}_{n-1}] = \infty \right\}$$

Observe that the first two Borel-Cantelli lemmas follow from this one.

#### Proof.

Define

$$X_n = \sum_{k=1}^n (1_{A_k} - E[1_{A_k} \mid \mathcal{F}_{k-1}])$$

(with  $X_0 = 0$ ). This is a martingale as

$$E[X_n - X_{n-1} \mid \mathcal{F}_{n-1}] = E[1_{A_n} - E[1_{A_n} \mid \mathcal{F}_{n-1}] \mid \mathcal{F}_{n-1}] = 0.$$

Clearly  $|X_n - X_{n-1}| \le 1$ . We apply the previous theorem, noting that on both C and D

$$\sum_{k=1}^{\infty} 1_{A_k} = \infty \Longleftrightarrow \sum_{k=1}^{\infty} E[1_{A_k} \mid \mathcal{F}_{k-1}] = \infty.$$



### Theorem (Three series theorem)

Let  $\xi_1, \xi_2, \ldots$  be independent r.v.'s. For A>0, let  $\eta_n=\xi_n \mathbf{1}_{\{|\xi_n|\leq A\}}$ . Consider the following three series (i) $\sum_{n=1}^{\infty}P(|\xi_n|\geq A)$ ; (ii) $\sum_{n=1}^{\infty}E\eta_n$ ; and (iii) $\sum_{n=1}^{\infty}\mathrm{Var}(\eta_n)$ . Then the following are equivalent:

- (a)  $\sum_{n=1}^{\infty} \xi_n$  converges a.s.;
- (b) (i), (ii), and (iii) converge for all A > 0; and
- (c) (i), (ii), and (iii) converge for some A > 0.

Recall that  $\{\sum_{n=1}^{\infty} \xi_n\}$  converges $\}$  is a tail event, so by the 0-1 law, the statement (a) is equivalent to  $P(\sum_{n=1}^{\infty} \xi_n\}$  converges) > 0.

**Example**. Assume that  $\epsilon_n$  are i.i.d., with  $P(\epsilon_n = \pm 1) = \frac{1}{2}$ . When

does 
$$S = \sum_{n=1}^{\infty} \epsilon_n \frac{1}{n^p}$$
 converge?

We only need to check when (iii) converges (with  $\xi_i$  instead of  $\eta_i$ ). This gives the condition

$$\sum_{p=1}^{\infty} \frac{1}{n^{2p}} < \infty \iff p > 1/2$$

If  $0 , the partial sums <math>S_n = \sum_{k=1}^n \epsilon_k \frac{1}{k^p}$  are a.s. dense in  $\mathbb{R}$ . This follows because  $S_n$  is a martingale with bounded increments.

When p > 1/2, we can determine the distribution of S by its characteristic function

$$E[e^{itS}] = E[e^{it\sum_n \epsilon_n n^{-\rho}}] = \prod_{n=1}^{\infty} E[e^{it\epsilon_n n^{-\rho}}] = \prod_{n=1}^{\infty} \cos(tn^{-\rho}).$$

#### Proof of Three series theorem.

- (c)  $\Longrightarrow$  (a)  $S_n = \sum_{k=1}^n (\eta_k E\eta_k)$  is a martingale with  $\langle S \rangle_n = \sum_{k=1}^n \mathrm{Var}(\eta_k)$ . By convergence of (iii),  $\langle S \rangle_\infty < \infty$ , and so  $S_n$  converges a.s. Then, as (ii) converges,  $\sum_{k=1}^\infty \eta_k$  converges a.s.
- By convergence of (i),  $P(\xi_k \neq \eta_k \text{ i.o.}) = 0$ , and so  $\sum_{k=1}^{\infty} \xi_k$  converges a.s.
- $(b) \Longrightarrow (c)$  Trivial.

#### Proof of Three series theorem, continued.

$$(a) \Longrightarrow (b)$$

Fix an A>0. If (i) diverges, then  $P(|\xi_k|>A \text{ i.o.})=1$  and so  $\sum \xi_k$  diverges a.s. So, (i) must converge and then  $P(\xi_k\neq \eta_k \text{ i.o.})=0$  so that  $\sum \eta_k$  converges a.s.

Now we apply the "symmetrization trick." Let  $\eta_k \stackrel{d}{=} \eta_k'$ , and construct the sequence  $(\eta_k')$  of independent r.v.'s, which is independent of  $(\eta_k)$ .

Then  $\sum_k \eta_k'$  also converges a.s. and then so does  $\sum_k (\eta_k - \eta_k')$ . Now  $S_n' = \sum_{k=1}^n (\eta_k - \eta_k')$  is a martingale with bounded increments, and so its bracket  $\langle S' \rangle_n = \sum_{k=1}^n \mathrm{Var}(\eta_k - \eta_k') = \sum_{k=1}^n 2\mathrm{Var}(\eta_k)$  must converge. It follows that  $\sum_{k=1}^\infty \mathrm{Var}(\eta_k) < \infty$ . So,  $\sum_{k=1}^n (\eta_k - E\eta_k)$  converges a.s., as its bracket converges. Thus,  $\sum_{k=1}^\infty E\eta_k$  converges.