On asymmetric simple exclusion process with periodic step Bernoulli initial condition

Craig A. Tracy\textsuperscript{1, a)} and Harold Widom\textsuperscript{2, b)}

\textsuperscript{1}Department of Mathematics, University of California, Davis, California 95616, USA
\textsuperscript{2}Department of Mathematics, University of California, Santa Cruz, California 95064, USA

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We consider the asymmetric simple exclusion process (ASEP) on the integers in which the initial density at a site (the probability that it is occupied) is given by a periodic function on the positive integers. (When the function is constant, this is the step Bernoulli initial condition.) Starting with a result in earlier work, we find a formula for the probability distribution for a given particle at a given time which is a sum over positive integers $k$ of integrals of order $k$.

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I. INTRODUCTION

In the asymmetric simple exclusion process (ASEP) on the integers $\mathbb{Z}$, a particle waits exponential time, then moves to the right with probability $p$ if that site is unoccupied (or else stays put) or to the left with probability $q = 1 - p$ if that site is unoccupied (or else stays put). A formula was derived in Ref. 14 for $P_y(x_\ell(t) \leq x)$, the probability distribution function for $x_\ell(t)$, the position of the $\ell$th particle from the left at time $t$, given the initial configuration $Y = \{y_1 < y_2 < \cdots \}$. The set $Y$ may be finite or infinite on the right. The formula, given below in (1), is a sum over all (finite) subsets $S$ of $Y$ of integrals of order $|S|$.

For step initial condition, when $Y = \mathbb{Z}^+$ (the positive integers), three pleasant things happen:

(i) For each $k$ one can sum over all $S$ with $|S| = k$, thus replacing a sum over all finite $S \subset \mathbb{Z}^+$ by a sum over $k$.

(ii) There is a combinatorial identity that replaces the integrands by simpler ones which are symmetric in the integration variables.

(iii) The resulting $k$-dimensional integrals are coefficients in the expansion of a Fredholm determinant.

This led to a representation of the probability distribution in terms of a Fredholm determinant\textsuperscript{15, 16}, which was amenable to asymptotic analysis. In particular, Ref. 16 establishes Kardar-Parisi-Zhang (KPZ) Universality for ASEP. In Refs. 1, 11, and 12, the starting point for their analysis of the KPZ equation\textsuperscript{8} is this determinant representation in ASEP. The asymptotic analysis in the weakly asymmetric limit of ASEP, called WASEP, establishes the KPZ Universality for the KPZ equation\textsuperscript{1, 11, 12}. We refer the reader to the reviews\textsuperscript{5–7, 13} for a broader discussion of this field.

With step Bernoulli initial condition, $Y$ is not deterministic, but each site in $\mathbb{Z}^+$ is, independently of the others, initially occupied with probability $\rho$. In this case also we have (i)–(iii), and this made asymptotic analysis possible\textsuperscript{4, 17}.

In another direction, Lee\textsuperscript{9} considered the cases when $Y = m\mathbb{Z}^+$, and found that (i) and (ii) hold when $m = 2$, but only (i) holds when $m > 2$. Even for $m = 2$, (iii) does not seem to hold since the $k$-dimensional integrals are not, or not obviously, related to a determinant or Pfaffian. (Earlier work on the totally asymmetric simple exclusion process (TASEP) with initial condition $2\mathbb{Z}^+$ or $2\mathbb{Z}$ is in Refs. 2, 3, and 10.)

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\[ a) \text{Electronic mail: tracy@math.ucdavis.edu.} \]

\[ b) \text{Electronic mail: widom@ucsc.edu.} \]
Nevertheless, it seems worthwhile to consider a generalization which we call \textit{periodic step Bernoulli initial condition}. Here we have an \(m\)-periodic function \(n \rightarrow \rho_n\) from \(\mathbb{Z}^+\) to \([0, 1]\), and assume that initially a site \(n \in \mathbb{Z}^+\) is occupied with probability \(\rho_n\). (Initially, sites \(n \not\in \mathbb{Z}^+\) are empty.) The most that can be expected in this generality is that (i) holds, and we shall show that it does.

Here are the formulas alluded to and the result of the present paper. We assume \(q \neq 0\), set \(\tau = p/q\), and recall that the \(\tau\)-binomial coefficient \(\binom{N}{\ell}_\tau\) is defined as

\[
\frac{(1 - \tau^N)(1 - \tau^{N-1})\cdots(1 - \tau^{N-\ell+1})}{(1 - \tau)(1 - \tau^2)\cdots(1 - \tau^\ell)}.
\]

We define

\[
\varepsilon(\xi) = p\xi^{-1} + q\xi - 1, \quad f(\xi_i, \xi_j) = \frac{\xi_j - \xi_i}{p + q\xi_j\xi_j - \xi_i},
\]

and then

\[
I(k, \xi) = I(k, \xi_1, \ldots, \xi_k) = \prod_{i<j} f(\xi_i, \xi_j) \prod_{i} \xi_i^{\varepsilon(\xi_i)} \frac{e^{\tau(\xi_i)^\ell}}{1 - \xi_i}.
\]

All indices in the products run from 1 to \(k\). Notice that \(I(k, \xi)\) depends on \(x\) and \(t\), although they are not displayed in the notations.

Finally, given two sets of integers \(U\) and \(V\), we define

\[
\sigma(U, V) = \#\{(u, v) : u \in U, v \in V, \text{ and } u \geq v\}.
\]

Theorem 5.2 of Ref. 14, in different notation, is the formula

\[
\mathbb{P}_Y(x_i(t) \leq x) = \sum_{S \subset Y} c_{\ell, k} \varepsilon^{(S, Y)} \prod_{i=1}^k I(k, \xi) \prod_{i=1}^k \xi_i^{-\varepsilon},
\]

where \(k = |S|\) and

\[
S = \{s_1, \ldots, s_k\}, \quad (s_1 < s_2 < \cdots < s_k).
\]

\[
c_{\ell, k} = (-1)^{\ell} q^{k(k-1)/2} \tau^{\ell(k-\ell)/2} \left\lfloor \frac{k - 1}{\ell - 1} \right\rfloor.
\]

The sum is taken over all (finite) subsets \(S\) of \(Y\) with \(|S| \geq \ell\). The contour \(C_R\) is the circle with center 0 and radius \(R\), which is assumed so large that the denominators \(p + q\xi_j\xi_j - \xi_i\) are nonzero on and outside the contours. (All integrals are given the factor \(1/2\pi i\).)

When \(Y\) is not deterministic, the probability \(\mathbb{P}(x_i(t) \leq x)\) is obtained by averaging the right side of (1) over all \(Y\). The focus of the present note is (i), the computation of the sum over all \(S \subset Y\) with \(|S| = k\), and then the average over all initial configurations \(Y \subset \mathbb{Z}^+\). What will be important here are only those ingredients of (1) that depend on \(Y\) and \(S\). These combine as

\[
\varepsilon^{(S, Y)} \prod_{i=1}^k \xi_i^{-\varepsilon}.
\]

For step initial condition, the sum over (2) over all \(S\) equals [p. 838 of Ref. 14]

\[
\tau^{k(k+1)/2} \prod_{i=1}^k \frac{1}{\xi_i \cdots \xi_k - \tau^{k+1}}.
\]

For step Bernoulli initial condition, the sum over \(S \subset Y\) followed by the average over initial configurations \(Y\) is [p. 830 of Ref. 17]

\[
\tau^{k(k+1)/2} \prod_{i=1}^k \frac{\rho}{\xi_i \cdots \xi_k - 1 + \rho - \tau^{k+1}\rho}.
\]
For $Y = m\mathbb{Z}^+$, the sum of (2) over $S$ equals
\[ \tau^{k(k+1)/2} \prod_{i=1}^{k} \frac{1}{(\xi_i \cdots \xi_k)^m - \tau^{k-i+1}}. \] (4)

To state the formula we derive here for periodic step Bernoulli initial condition, we define
\[ \varphi(i, n) = 1 - \rho_n + \rho_n \tau^{k-i+1} \]
and $m \times m$ matrices $A_i$ with entries
\[ A_{i, \mu, \nu} = \frac{1}{(\xi_i \cdots \xi_k)^m - \prod_{n=1}^{m} \varphi(i, n)} \times \begin{cases} (\xi_i \cdots \xi_k)^m \xi_i^{-\nu} \rho_i \prod_{n=\mu+1}^{-1} \varphi(i, n), & \text{if } \mu < \nu, \\ \xi_i^{-\nu} \rho_i \prod_{n=\mu+1}^{-1} \varphi(i, n), & \text{if } \mu \geq \nu. \end{cases} \] (5)

The indices $\mu$ and $\nu$ run from 0 to $m-1$.

The result is that the average over $Y$ of the sum of (2) over $S \subset Y$ with $|S| = k$ is equal to $\tau^{k(k+1)/2}$ times
\[ (1 0 \cdots 0) \prod_{i=1}^{k} A_i \begin{pmatrix} 1 \\ \vdots \\ 1 \end{pmatrix}. \] (6)

This is, in words, the sum of the entries in the top row of the matrix product. In the product, matrices with lower index $i$ are on the left. To recapitulate:

**Theorem:** For periodic step Bernoulli initial condition, the probability $\mathbb{P}(x_i(t) \leq x)$ is given by the right side of (1) with $\sum_{S \subset Y}$ replaced by $\sum_{k \geq i}$, and the product (2) replaced by $\tau^{k(k+1)/2}$ times (6).

In the simplest special cases, $\rho_n = \rho$ when $n \equiv v \pmod{m}$, and $\rho_n = 0$ otherwise. If we take $0 \leq v < m$, then (6) is equal to
\[ A_{1,0,v} \prod_{i=2}^{k} A_{i,v,v}, \]
which we find is equal to
\[ \prod_{i=1}^{k} (\xi_i \cdots \xi_k)^m - (1 - \rho + \rho \tau^{k-i+1}) \]
when $v = 0$, and
\[ (\xi_1 \cdots \xi_k)^m \prod_{i=1}^{k} (\xi_i \cdots \xi_k)^m - (1 - \rho + \rho \tau^{k-i+1}) \]
otherwise. Alternatively, it is given by just the last formula if instead we take $0 < v \leq m$.

In the course of the proof we obtain an analogous statement for what we may call the general step Bernoulli initial condition, in which the function $\rho : \mathbb{Z}^+ \to [0, 1]$ is arbitrary. We define $B_i$ to be the semi-infinite matrix with entries
\[ B_{i, \mu, \nu} = \begin{cases} \xi_i^{-\nu} \rho_i \prod_{n=\mu+1}^{v} \varphi(i, n), & \text{if } \mu < \nu, \\ 0, & \text{if } \mu \geq \nu. \end{cases} \] (7)
the indices satisfying $0 \leq \mu, \nu < \infty$. We shall see that the average over $Y$ of the sum of (2) over $S \subset Y$ with $|S| = k$ is then equal to $\tau_{k+1}^{2k+1}$ times

$$\left(1 0 0 \cdots \right) \prod_{i=1}^{k} B_{i} \begin{pmatrix} 1 \\ 1 \\ \vdots \end{pmatrix}. \tag{8}$$

Observe that this is what is obtained by formally letting $m \to \infty$ in the periodic case. Observe also that if $\rho$ takes only the values 0 or 1, we are back in the case of deterministic initial configuration, and what we obtain can be seen to be just a restatement of (1).

II. PROOF OF THE THEOREM

We assume first that $\rho : \mathbb{Z}^+ \to [0, 1]$ is arbitrary and want to compute the average over $Y$ of the sum of (2) over $S \subset Y$ with $|S| = k$. As in Ref. 17, we do this in the opposite order. We first take a fixed $S$ and average over all $Y \supset S$. (Later we will take the sum over $S$.) Also, as in Ref. 17, we may assume that $Y \subset [1, N]$, and at the end, we let $N \to \infty$. (In fact all we shall use is that $N \geq s_{k}$.)

A. The average over $Y$

The part of (2) that depends on $Y$ is $\tau^{\sigma(S, Y)}$, while the probability of an initial configuration $Y$ is

$$\prod_{n \in Y} \rho_{n} \prod_{n \in [1, N] \setminus Y} (1 - \rho_{n}),$$

so we have to compute

$$\sum_{S \subset Y \subset [1, N]} \tau^{\sigma(S, Y)} \prod_{n \in Y} \rho_{n} \prod_{n \in [1, N] \setminus Y} (1 - \rho_{n}).$$

Any element of $Y$ which is larger than $s_{k}$ does not affect $\sigma(S, Y)$, so this sum may be written as a product of sums

$$\left(\sum_{S \subset Y \subset [1, s_{k}]} \tau^{\sigma(S, Y)} \prod_{n \in Y} \rho_{n} \prod_{n \in [1, s_{k}] \setminus Y} (1 - \rho_{n})\right) \cdot \left(\sum_{Y \subset [s_{k}, N]} \prod_{n \in Y} \rho_{n} \prod_{n \in [s_{k}, N] \setminus Y} (1 - \rho_{n})\right). \tag{9}$$

The second factor equals 1, since it is

$$\prod_{n \in [s_{k}, N]} (\rho_{n} + (1 - \rho_{n})), $$

when written as a sum of products.

It remains to consider the first factor in (9). We use the notation

$$Y_{i} = (s_{i-1}, s_{i}) \cap Y, \quad (s_{0} = 0),$$

so $Y$ is the disjoint union $S \cup Y_{1} \cup \ldots \cup Y_{k}$. Since the number of elements of $S$ greater than or equal to $s_{i}$ is $k - i + 1$, we have

$$\sigma(S, Y) = \sigma(S, S) + \sum_{i=1}^{k} \sigma(S, Y_{i}) = k(k + 1)/2 + \sum_{i=1}^{k} (k - i + 1) |Y_{i}|.$$
Therefore, the first factor in (9) equals
\[
\tau^{k(k+1)/2} \prod_{n \in S} \rho_n \prod_{i=1}^{k} \left( \sum_{Y_i \subset (s_i-1, s_i)} \tau^{(k-i+1)|Y_i|} \prod_{n \in Y_i} \rho_n \prod_{n \in (s_i-1, s_i) \setminus Y_i} (1 - \rho_n) \right)
\]
\[
= \tau^{k(k+1)/2} \prod_{n \in S} \rho_n \prod_{i=1}^{k} \left( \sum_{Y_i \subset (s_i-1, s_i)} \prod_{n \in Y_i} (\tau^{k-i+1} \rho_n) \prod_{n \in (s_i-1, s_i) \setminus Y_i} (1 - \rho_n) \right).
\]

The sum over all \( Y_j \subset (s_{j-1}, s_j) \) is
\[
\prod_{n \in (s_{j-1}, s_j)} \varphi(i, n) = \prod_{n \in (s_{j-1}, s_j)} (\tau^{k-i+1} \rho_n + (1 - \rho_n)),
\]
as is seen by writing this as a sum of products. Hence (9) equals \( \tau^{k(k+1)/2} \) times
\[
\prod_{n \in S} \rho_n \prod_{i=1}^{k} \prod_{n \in (s_{i-1}, s_i)} \varphi(i, n).
\]

(10)

B. The sum over \( S \): general \( \rho \)

This gives the average of \( \tau^\sigma(S, Y) \) over all \( Y \supset S \). Now we are to multiply this by \( \prod_i \xi_i^{-s_i} \), the second factor in (2), and sum over all \( S \) with \( |S| = k \). Since \( S = \{s_1, \ldots, s_k\} \), the sum over all \( S \) equals
\[
\sum_{0 < s_1 < \ldots < s_k} \prod_{i=1}^{k} \left( \rho_{n_i} \xi_i^{-s_i} \prod_{n \in (s_{i-1}, s_i)} \varphi(i, n) \right).
\]

(11)

In terms of the entries of the matrices \( B_i \) given by (7), this equals
\[
\sum_{s_1, \ldots, s_k=0}^{\infty} B_{1, 0, s_1} B_{2, s_1, s_2} \cdots B_{k, s_{k-1}, s_k},
\]
which equals (8), the sum of the entries of the top row of the matrix product. (We used here that \( B_{i, \mu, v} = 0 \) unless \( \mu < v \).)

C. The sum over \( S \): periodic \( \rho \)

We shall rewrite the sum (11) using the periodicity of \( \rho \). There are unique representations
\[
s_i = u_i m + v_i, \quad \text{where} \quad u_i \geq 0 \quad \text{and} \quad 0 \leq v_i < m.
\]

We have \( s_{i-1} < s_i \) if and only if the following hold:
\[
\begin{align*}
\text{if} \quad v_{i-1} < v_i, & \quad \text{then} \quad u_{i-1} \leq u_i, \\
\text{if} \quad v_{i-1} \geq v_i, & \quad \text{then} \quad u_{i-1} < u_i.
\end{align*}
\]

If we set \( u_i = t_1 + \cdots + t_i \), so that \( s_i = (t_1 + \cdots + t_i) m + v_i \), then the above becomes
\[
\begin{align*}
t_i & \geq 0, \quad \text{if} \quad v_{i-1} < v_i, \\
t_i & > 0, \quad \text{if} \quad v_{i-1} \geq v_i.
\end{align*}
\]

(12)

With this notation the factor \( \xi_i^{-s_i} \) in (11) becomes
\[
\xi_i^{-(t_1 + \cdots + t_i) m - v_i}.
\]
the factor $\rho_i$ becomes $\rho_v$ by the periodicity of $\rho$, and by the periodicity of $\varphi(i, n)$ in $n$ the last factor becomes
\[
\prod_{n=v_i-1+1}^{v_i+m-1} \varphi(i, n) \cdot \left( \prod_{n=1}^{m} \varphi(i, n) \right)^{v_i-1}, \quad \text{if } v_i-1 < v_i,
\]
\[
\prod_{n=v_i-1+1}^{v_i+m-1} \varphi(i, n) \cdot \left( \prod_{n=1}^{m} \varphi(i, n) \right)^{v_i-1-1}, \quad \text{if } v_i-1 \geq v_i.
\]
If we fix the $v_i$ and first take the sum in (11) over the $t_i$ according to (12), we obtain
\[
\frac{1}{(\xi_1 \cdots \xi_k)^m} \prod_{i=1}^{m} \varphi(i, n) \times \prod_{n=v_i-1+1}^{v_i+m-1} \varphi(i, n), \quad \text{if } v_i-1 < v_i,
\]
\[
\prod_{n=v_i-1+1}^{v_i+m-1} \varphi(i, n), \quad \text{if } v_i-1 \geq v_i.
\]
Then the product of this over $i$ is to be summed over all $v_i$ satisfying $0 \leq v_i < m$. In terms of the entries of the matrices $A_i$ given by (5), the sum equals
\[
\sum_{v_1, \ldots, v_k = 0}^{m-1} A_{1,0} A_{2,v_1} A_{3,v_2} \cdots A_{k,v_{k-1},v_k},
\]
which equals (6). This completes the proof of the theorem.

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1 Amir, G., Corwin, I., and Quastel, J., “Probability distribution of the free energy of the continuum directed random polymer in 1 + 1 dimensions” Commun. Pure Appl. Math. 64, 466 (2011).